

# The Intertwining of Credit and Banking Fragility\*

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## Abstract

While the literature has provided evidence of the predictive power of credit for financial and banking crises, this paper aims to investigate the grounds of this link by assessing the interrelationships between credit and banking fragility. The main identification assumption represents credit and banking fragility as a system of simultaneous joint data generating processes whose error terms are correlated. We test the null hypotheses that credit positively affects banking fragility -a vulnerability effect- and that banking fragility has a negative effect on credit -a trauma effect-. We use Seemingly Unrelated Regressions and 3SLS on a panel of European Union (EU) countries from 1998 to 2012 and control for the financial and macroeconomic environment. We find a positive effect of credit on banking fragility in the EU as a whole, in the Eurozone, in the core of the EU but not at its periphery, and a negative effect of banking fragility on credit in all samples.

**Keywords:** Credit growth, banking fragility, non-performing loans, SUR model.

**JEL Classification:** E44; G20.

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## 1. Introduction

The objective of this paper is to link two strands of the literature. The first literature examines the nature of financial and banking crises and their determining factors (e.g. Allen and Gale, 2009, Barro, 2009, or Almunia et al, 2010). In this vein, Schularick and Taylor (2012) and Aikman et al. (2015) provide evidence, over a long era and for a large sample of countries, of the predictive power of credit for financial crises. The second one investigates the consequences of financial and banking crises on the subsequent recovery. Some papers (e.g. Brunnermeier and Pedersen, 2009; Geanakoplos, 2010; Shleifer and Vishny, 2011) focus on the behaviour of the banking sector in the aftermath of such crises. In this paper we explore the interrelationships of credit and banking fragility in the European Union (EU). Three reasons motivate this study. First, the global financial crisis has shed light on the intertwining between the growth of the banking and financial sectors, financial deregulation and banking fragility (e.g. Gorton and Metrick, 2012). Second, the EU has adopted a banking union which gives the European Central Bank (ECB) a role of prudential supervisor for most banks in the EU. The ECB is *de facto* in charge of monitoring credit and bank stability. Third, although the determinants of credit, measured as the ratio of domestic credit to the private sector to GDP, and the determinants of banking fragility, such as non-performing loans (NPL), have been investigated separately in the empirical literature, their cross-relationships have not been yet to our knowledge.

The use of the share of NPL to gross loans as a proxy for banking fragility is motivated by the outcomes of Cihak and Schaeck (2010). They find that the contemporaneous ratio of NPL to total loans provides relevant warning signals for systemic banking crisis. High levels of NPL constrain bank capital that could otherwise be used to increase lending. Aoki and Nikolov (2015) also show that the real effects of bubbles crucially depend on the identity of the bubble holder. Bubbles held by banks lead to a larger boom-bust cycle in credit and output compared to bubbles held by ordinary savers. High levels of NPL not only raise financing costs for small and medium enterprises, but also trigger financial crisis and have devastating real effects.

We limit our investigation to the period 1998-2012 for which banking, macroeconomic, and market data are available for most of the EU countries. Figure 1 shows a scatter plot of NPL to total gross loans and credit to GDP. The relationship is unclear and the unconditional correlation is -0.23. In contrast, the contribution of this paper is to assess the conditional correlation between credit and banking fragility and to single out the effect of each of these two variables on the other. We impose a panel structure on data and control for time and country fixed effects, as well as financial and macroeconomic environment. The latter encompasses potential determinants of bank credit, as shown in the literature: GDP growth, inflation, and trade openness, and potential determinants of banking fragility: long-term real interest rates, taxes on business, a financial regulation index and market capitalisation.

We test the following two null hypotheses: (i) there is a positive effect of credit on banking fragility labelled a “vulnerability effect” and (ii) there is a negative effect of banking fragility on credit that we label a “trauma effect”. The first hypothesis stems from the increasing fragility and risks of marginal loans, whereas the second results from the potential deleveraging and reduced risk-taking of banks following a period of banking fragility.

While estimating the link between credit and banking fragility, we are confronted to two types of endogenous processes. The first is related to the joint determination of the two left-hand-side variables. Like price and quantity on a given market, credit and banking fragility

can be considered as the opposite sides of the same coin. To correct for their simultaneity, we represent credit and banking fragility as a system of simultaneous joint data generating processes estimated with Seemingly Unrelated Regressions (SUR) which takes into account that contemporaneous error terms are correlated and provides more efficient estimates than OLS. The second type of endogeneity relates to the right-hand-side variables and to the estimation of their causal effect. A potential omitted variable bias or reverse causality would make these variables and the error term correlated. This second type of endogeneity is handled with instrumental variables. We perform a three-stage least squares (3SLS) estimation which enables to combine the system estimation of SUR with the instrumental-variable method of 2SLS.

Despite the negative correlation between credit and banking fragility, presented in Figure 1, we find a positive causal effect of the level of credit to GDP on the share of NPL, and a negative causal effect of NPL on credit. These results are robust to using the growth rate of credit, alternative banking fragility variables, the introduction of government debt, to most EU subsamples, to non-linear specifications and to a 3-equation SUR model in which long-term interest rates are also considered endogenous. More precisely, we find the existence of a vulnerability effect in the EU as a whole, in the Eurozone, in the core of the EU but not at its periphery. We attribute the difference between the core and the periphery to their different stages of financial development. We also find evidence of non-linearities between the two main variables. NPL have a non-linear effect on credit to GDP depending on the level of credit to GDP, while the effect of credit to GDP on NPL –the vulnerability effect– depends on the level of credit to GDP and is time contingent: this effect kicks-in during crisis times.

The rest of this paper is organized as follows. Section 2 reviews the literature. Section 3 describes the model, the empirical strategy and the hypotheses. Section 4 presents the data. Section 5 discusses the results. Section 6 concludes.

## **2. Determinants of credit and NPL in the literature**

This paper relies on the literature about credit and its determinants. After the extension of the IS/LM model to banks by Bernanke and Blinder (1988), this literature expanded on the analysis of monetary policy channels of transmission, whereas the bulk of empirical papers about credit devoted attention to its impact on economic growth (see Ang, 2008, for a survey). Only a few papers investigated credit determinants. Following Goodhart (1995), Hofmann (2004) shows that shocks to property prices could explain the persistence in financial cycles. In the vein of Kashyap and Stein (1995), Ashcraft (2006) studies the lending channel in the US economy and uses the affiliation with multibank holding companies to proxy financial constraints across banks. He finds that annual loan growth of affiliated banks is less sensitive to federal funds rates than non-affiliated banks. Altunbas et al. (2009) extend Ashcraft (2006)'s empirical model to the securitisation activities of European banks. They show that securitisation helps banks circumvent the impact of monetary policy. They also relate the growth of bank loans to bank risks and estimate the link between credit and loan loss provisions. The latter has the significant negative expected sign *vis-à-vis* the former. Cottarelli et al. (2005) study the credit growth in Central and Eastern European countries (CEECs) and test whether it could be attributed to a structural change of financial deepening. Their list of credit determinants includes public debt to GDP ratio, GDP per capita, an indicator of high inflation, an indicator of financial liberalization, and different institutional characteristics like accounting standards, legal origins and bank entry requirements. Except for the latter, all variables have the significant expected sign. Aisen and Franken (2010) explain real credit growth in 83 countries, with a distinction between, first, variables of

economic performance, external shocks and policy stance; second, local characteristics of the credit market (like size, integration, and openness); and third, bank characteristics *per se* (like share of public ownership, bank leverage, and bank return on equity). GDP growth and changes in money market rate are the significant ones. Chinn and Ito (2006) discuss the role of capital controls and institutions on credit, thus questioning the relationship between financial openness and financial development. Dell’Ariccia et al. (2016) identify three factors that trigger credit booms in 170 countries over the period 1970–2010: financial reforms and strong economic growth. At a micro level, Aiyar et al. (2014) investigate the supply of credit and its linkages with (and leakages towards) credit substitution channels via foreign affiliates and branches to comply with macro-prudential measures.

The literature on banking fragility and its determinants has developed along two different lines of reasoning.<sup>1</sup> The first one assumes that capitalism is intrinsically unstable (Minsky, 1995) and leads to leverage and credit booms and busts. The second one sticks to a general equilibrium approach and assumes that banking fragility is caused by financial frictions (due to asymmetric information), hence by financial shocks and their propagation to the rest of the economy (Calomiris, 1995; Mishkin, 1999). The share of NPL in bank balance sheets has been shown to trigger the onset of a banking crisis (Reinhart and Rogoff, 2011). Louzis et al. (2012) study the macroeconomic and bank-specific determinants of NPL in Greece, and find that they mostly respond to GDP, unemployment, interest rates and public debt. Finally, Ruiz-Porrás (2009) assess the effects of financial structure and financial development on banking fragility while Gropp, Vesala and Vulpes (2006) analyse the properties of distances-to-default and bond spreads as leading indicators of banking fragility.

### 3. Empirical strategy and null hypotheses

While assessing the link between credit and banking fragility, we face the issue of their potential endogeneity. One solution, and this is the main identification assumption of this paper, consists in thinking the problem not in a single-equation space, but as a system of simultaneous equations that jointly determine both dependent variables. The two equations are therefore mechanically related as the contemporaneous errors associated with each dependent variable are correlated, which seems a reasonable assumption for these two data processes. The most basic form of joint-system estimation is Seemingly Unrelated Regressions (SUR), also called Zellner (1962)-efficient regressions, using feasible generalised least-squares (FGLS). When the two equations do not have the same set of explanatory variables and are not nested, it leads to more efficient estimates than estimating each individual equation separately with OLS because it takes into account the correlation between the error terms and therefore adds information on the error structure. Generally, the coefficients are only slightly different, but the standard errors are uniformly larger.

We estimate simultaneously the cross-effects of credit and banking fragility using the following model, in which we assess the contribution of our variables of interest above and beyond contemporaneous financial and macro controls and past information captured by the lagged value of the dependent variables:

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<sup>1</sup> Other measures of banking fragility than NPL have been proposed. Loayza and Ranciere (2006) measure it as the standard deviation of the growth rate of the private credit to GDP ratio over non-overlapping 5-year averages. The ECB has developed a Composite Indicator of Systemic Stress (CISS) for the euro area. The International Monetary Fund (IMF) developed financial soundness indicators. At the micro level, several authors capture financial stability in the banking sector through the Z-score (Uhde and Heimeshoff, 2009; Fink et al., 2009), which measures the probability of default for a bank or a banking system.

$$\begin{cases} F_{i,t} = \alpha_F + \beta_F F_{i,t-1} + \beta_{FC} C_{i,t} + \beta_{FC'} C_{i,t-1} + \beta_{FX} X_{i,t} + \beta_{FZ} Z_{i,t} + \varepsilon_{F,t} \\ C_{i,t} = \alpha_C + \beta_C C_{i,t-1} + \beta_{CF} F_{i,t} + \beta_{CF'} F_{i,t-1} + \beta_{CX} X_{i,t} + \beta_{CZ} Z_{i,t} + \varepsilon_{C,t} \end{cases} \quad (1)$$

where  $F_{i,t}$  is the banking fragility variable for country  $i$ ,  $C_{i,t}$  is the credit variable,  $X_{i,t}$  is a vector of financial controls, namely long-term real interest rates, stock market capitalisation, taxes and a financial regulation variable, and  $Z_{i,t}$  includes country and time fixed effects and the macroeconomic environment, namely real GDP, inflation and trade openness. Given the annual frequency of the data and the fact that the length between a loan disbursement and its possible classification as NPL is at least 90 days, the emission of a credit line and its reclassification as a NPL may happen during the same year, so we include a potential contemporaneous relationship between credit and banking fragility. Using this model, we test two hypotheses:

*Hypothesis n°1*: there is a positive effect of credit on banking fragility labelled a “vulnerability effect”.

This vulnerability effect stems from the increasing fragility and risks of marginal loans. This effect also arises from the dependence of loan-loss provisioning on the evolution of bank lending. Pool et al. (2015) show that banks reduce their loan-loss provisioning as a percentage of their total assets when bank lending increases, and therefore take on more risks. Gourinchas and Obstfeld (2012), Schularick and Taylor (2012) and Aikman et al. (2015) show that rapid domestic credit expansion is a robust indicator of financial crises. One could expect a U-shaped relationship between credit and NPL. Until a threshold, credit will help develop an efficient market for loans while the marginal utility of bank loans will be positive. However, once a threshold is reached, the risk of marginal loans increases. One could also expect the occurrence of a relationship that takes a convex form between credit and NPL: the risk of marginal loans increases disproportionately with the supply of loans. We therefore test for possible non-linearities of this relationship.

*Hypothesis n°2*: there is a negative effect of banking fragility on credit that we label a “trauma effect”.

This effect results from the potential deleveraging and reduced risk-taking of banks following a period of banking fragility. This is suggested by Adrian and Shin (2010, 2014) who theoretically document the procyclicality of the leverage of financial intermediaries. They show that financial intermediaries maintain a constant probability of default to shifts in the outcome distribution so it implies substantial deleveraging during downturns. This procyclicality may have been reinforced by regulatory measures. This hypothesis also relies on theoretical mechanisms that have been put forward by Brunnermeier and Pedersen (2009), Geanakoplos (2010) and Shleifer and Vishny (2011). Brei and Gambacorta (2016) show that the risk-weighted regulatory capital ratio of Basel III is less procyclical than the previous liquidity ratio, that was mandatory during our period of analysis. Similarly to the first hypothesis, one can expect non-linearities in the effect of banking fragility on credit: the deeper the crisis, the stronger the deleveraging and the negative effect on credit supply.

We include financial variables in the regression that could impinge on the relationships between credit and banking fragility.<sup>2</sup> We expect a negative effect of long-term real interest

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<sup>2</sup> Another potentially interesting variable would have been the degree of securitization, enabling to have credit to GDP and NPL corrected for securitization, so capturing all loans issued and not only those still on banks' balance sheet. Unfortunately, to our knowledge, data are not available for our sample.

rates (measuring financing costs) on credit. We assume that credit demand decreases and credit supply increases with interest rates. Fase (1995) reports results on credit for the Netherlands using nominal long-term interest rates. Alternatively, we focus on real long-term interest rates. We expect a positive correlation between the long-term real interest rate and banking fragility: the latter materializes after real interest rates go up, hence weakening debtors' positions. We expect a positive link between taxes and credit and between taxes and banking fragility. Following Keen and De Mooj (2012) and De Mooj, Keen and Orihara (2013), the corporate tax would violate the Modigliani-Miller theorem in the case of banking institutions: the high corporate tax induces recourse to borrowing (debt) to grasp the full benefit of interest payments' deduction, at the expense of equity. We expect a negative link between stock market capitalisation and credit which may capture a substitution effect between direct finance and bank intermediation. This may in turn induce a negative correlation between stock market capitalisation and banking fragility. Finally, we control for the existence of a positive link between financial deregulation and credit and a positive link between financial deregulation and banking fragility as deregulation may increase risk-taking. Chinn and Ito (2006) report a positive relationship between financial openness and financial development whereas Tressel and Detragiache (2008) show that financial liberalisation has a limited impact on financial development. Kaminsky and Schmukler (2008) show that financial liberalisation generates banking fragility in the short run.

In addition, we control for the effect of macroeconomic variables like the GDP growth rate, the inflation rate, and trade openness on credit and financial stability. Hofmann (2004) shows that a shock to real GDP can increase credit, e.g. in Germany, Ireland or Finland; or it can have no effect, e.g. in the USA, UK and Japan. Louzis et al. (2012) report a negative impact of GDP growth on NPL. Finally, Gozgor (2014) provides evidence of a positive link between trade openness and credit.

Two other issues, related to the onset of the global financial crisis and its European sequel, the sovereign-debt crisis, require some attention. First, the crisis has revealed the divergence between the Eurozone and the late newcomers in the EU, where the former have benefited from financial deepening for decades whereas the latter are in a process of financial development. The crisis has also revealed the gap between a core of EU countries and the periphery. These regional features may impinge on the relationship between credit and banking fragility and require a specific investigation. Second, growing public debts may affect credit demand and crowd out some investments as well as it may deteriorate the balance sheets of banks and thus modify credit supply and increase risks in the banking and financial system. Therefore, we test the potential effects coming from fiscal variables by introducing government debt.

## **4. Data**

### ***4.1. Dependent variables***

We measure credit with the level, or alternatively the growth rate, of the ratio of domestic credit to the private sector by deposit money banks and other financial institutions to GDP (in %) computed from the World Bank Global Financial Development Database (GFDD). We also use the deposit money banks' assets to GDP (%) as another measure of bank deepening. For the stock market view, we substitute credit to GDP by the turnover ratio (see Beck and

Levine, 2004). Banking fragility is captured with an aggregate prudential ratio: the ratio of NPL to gross loans.<sup>3</sup> For the stock market view, we use a stock price volatility variable.

#### *4.2. Explanatory variables*

GDP growth, the inflation rate and trade openness are included to control for the macroeconomic environment. We also include financial variables to control for factors that could affect the two variables of interest. Credit costs are captured by long-term real interest rates. The substitution effect between direct and indirect finance is tested with the stock market capitalisation or with the stock market turnover ratio. We assess the link between credit, banking fragility and taxes by using different measures of tax policies. Our benchmark measure is cyclically adjusted direct taxes on business. We also examine alternatively the ratio of total direct taxes to GDP, the ratio of capital taxes to GDP, and the ratio of cyclically adjusted taxes on production and imports to GDP. On the fiscal side, we consider the ratio of gross public debt to GDP. Finally, to isolate the effect of deregulation, we include an index of financial reform, or alternatively the level of bank regulatory capital to risk-weighted assets. All variables are described in Table A in the Appendix and descriptive statistics are presented in Table B.

#### *4.3. Subsample definitions*

There have been important evolutions in financial institutions due to liberalisation, innovation and globalisation, which have made differences between financial systems central to their analysis (Djankov et al., 2003). One important contribution in that respect is Bruno et al. (2012) who analyse the heterogeneity of financial systems through the lens of asset allocation among OECD countries. To shed light on the heterogeneity of the relationship between financial stability and credit into the EU, we decompose the sample into several subsamples.

First, we distinguish the Eurozone (EZ), composed of the 12 first member states of the euro area, leaving aside Luxembourg where banking deepening is so strong as to make this small country an outlier.

Second, the sovereign debt crisis highlighted the fragmentation in the EU. We then disentangle member states that belong to the core of the EU and member states that are more at the periphery. This separation is based on the spread between the domestic long-term sovereign interest rates and the German long-term sovereign interest rate post-2007. We choose the value of 0.80% as a cut-off criterion. Consequently, Spain and Italy are included in the periphery of the EU whereas the UK is part of the core.<sup>4</sup> The differences in the variables of the core EU and the EU periphery suggest that our grouping is reasonable. On the one hand, NPL, taxes on business, inflation and growth are on average higher in the periphery than in the core. On the other hand, credits to GDP and market capitalization are on average higher in the core than in the periphery. For robustness purposes, we propose another sample (Core 2) to test whether the inclusion of countries in the core (such as Spain or Italy) would change the results.

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<sup>3</sup> A loan is classified as a NPL when the payments of interest and principal are past due by 90 days or more.

<sup>4</sup> Usually, the distinction between core and periphery countries is realised for Eurozone countries. Here we study the links between credit and bank fragility in the EU, so non-Eurozone countries of the EU are included.

Third, we analyse another sub-sample (Newcomers) based on the recent waves of European enlargement. The composition of these sub-samples is available in Table C in the appendix together with a comparison of the mean and standard deviation of the main variables for all countries, and all sub-samples (see Table D).

## 5. Results

### 5.1. Baseline

Starting with our first hypothesis of a vulnerability effect, Table 1 shows that credit is a positive and significant determinant of banking fragility. This is true with or without the controls, but their inclusion reduces the magnitude of the effects (parameter estimates of controls are shown in Table E in the Appendix). Following Schularick and Taylor (2012) and for sake of clarity, we report the sum of the credit to GDP coefficients and its corresponding standard error.<sup>5</sup> When including controls (column 2), the coefficient is equal to 0.22 and is significant at the 1% level. According to our second hypothesis of a trauma effect, Table 1 shows that banking fragility (measured by NPL) has a negative effect on credit to GDP.<sup>6</sup> This is true with or without the financial and macro controls and the coefficient is equal to -0.15 and significant at the 1% level. Since all variables have been standardized to a normal distribution, this means that a 1-standard-deviation increase in NPL (namely, an increase of 5 percentage points of the share of NPL) reduces credit to GDP by 8 percentage points (the equivalent of 0.15 standard-deviation of the series of credit to GDP).<sup>7</sup> In both cases, the contemporaneous value of credit to GDP or NPL is not significant and suggests the existence of a dynamic process in the build-up of vulnerability and trauma effects. The last column of Table 1 shows estimates of equation (1) when the level of credit to GDP is replaced by the growth rate of credit to GDP. The positive effect of credit on banking fragility (the vulnerability effect) and the negative effect of NPL on credit (the trauma effect) are both confirmed. This suggests that this is not only the level of credit that matters but also the rhythm at which credit expands. Comparably, banking fragility has a negative influence on both the level and growth rate of credit.

We also assess in Table 1 the potential non-linear relations between credit and banking fragility. We first introduce squared values of each variable of interest as an explanatory variable of the other (column 3). We find that NPL have the same linear effect on credit to GDP whatever the NPL level, while the effect of credit to GDP on NPL –the vulnerability effect– is larger for high values of the credit to GDP ratio. More precisely, the effect of credit to GDP is small ( $0.34 - 0.22$ , so  $0.12$ ) and non-significant at one s.d. below the mean (36%) of the credit to GDP distribution whereas the effect is  $0.56$  ( $0.34 + 0.22$ ) and significant at the 1% level at one s.d. above the mean (151%) of its distribution. Second, we look at the cross-effects of each variable on the other by introducing an interaction term of the lagged dependent variable with the variable of interest (column 4). The effect of NPL on credit to GDP depends on the level of credit to GDP, whereas the effect of credit to GDP on NPL does not depend on the level of the share of NPL. For low values of credit to GDP (around 36%), the effect of NPL on credit to GDP is -0.07 but non-significant, whereas for high values of credit to GDP (around 151%), the effect of NPL on credit to GDP is negative (-0.21) and

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<sup>5</sup> We also report the sum of the NPL coefficients for the second equation of the system.

<sup>6</sup> As a robustness test, we also introduced the deposit banks assets as measure of bank deepening and the size of bank's balance sheet. Results hold and are available from the authors upon request.

<sup>7</sup> Figure 1 suggests some potential outliers for NPL. For robustness purposes, we removed data points above 20%. The raw correlation is -0.18 in that case. Column 2 of Table 1 has been re-estimated using that sample. Coefficients and t-stats are similar. These estimates are available from the authors upon request.



significant at the 1% level. It suggests that credit generates additional vulnerabilities. Finally, we consider the time-contingency of the effect and we interact the variable of interest with a dummy for the crisis taking the value 0 before 2007 and 1 from 2007 (column 5). NPL increase from 4.5% before 2007 to 5.2% after (with the s.d. decreasing from 5.3% to 4.4%) while credit increases from 78% before 2007 to 126% after (with the s.d. increasing from 48% to 62%). The effect of NPL on credit to GDP has not been altered during the financial crisis (the marginal effect is not significant, and the overall effect after 2007 is -0.18 and significant at the 5% level), whereas the vulnerability effect appears to kick-in during crisis times rather than during good times (the marginal effect is 0.52 and the overall effect after 2007 is 0.64 and significant at the 1% level). Interestingly, the crisis does not have an impact by itself. High levels of credit to GDP together with the occurrence of the crisis fuel banking fragility.

Finally, we estimate a 3-equation SUR model which includes long-term interest rates as a third simultaneous variable. Although we have been interested so far in the relationship between credit and banking fragility with long-term interest rates included in the set of explanatory variables, one can view long-term interest rates as another variable whose determination is simultaneous to credit and banking fragility. Credit demand depends directly on interest rates and the evolution of interest rates can trigger loan defaults as the subprime crisis showed. Column 6 in Table 1 provides estimates of the equation for the two main variables of interest and shows that they are not modified by this assumption. For the sake of parsimony, we pursue the rest of the analysis with a 2-equation SUR model.<sup>8</sup>

## 5.2. Estimating causal effects

So far, we have jointly estimated a set of equations assuming that they have no endogenous regressors. However, it is likely that the different variables on the right-hand-side of equations are endogenous. Using three-stage least squares (3SLS or SUR-IV) enables to combine the system estimation of SUR with the instrumental variables method of 2SLS so as to get a consistent estimator of equations with endogenous regressors. The 3SLS estimator works in 3 steps: 1. we calculate fitted values of the endogenous variables based on the reduced-form regressions on the exogenous variables as in 2SLS, 2. we estimate the individual equations by 2SLS, using their fitted values in place of the endogenous regressors, 3. we estimate the system of equations jointly by Generalized Least Squares.

Identification depends on two main assumptions: the instrument does not itself appear in the equation, and the instrument does appear in another equation that influences the endogenous regressor. This means that there needs to be one omitted exogenous variable for each included endogenous variable. There are two ways to assess the relevance of our instrumental variables. They should explain a significant share of the variation in the endogenous regressor, and they should be exogenous to the dependent variables, or in other words, they should not be correlated with the dependent variables except through their effects on the endogenous regressors. To check for the relevance of the instrumental variables, we provide the  $R^2$  of the regression of the 3SLS residuals on the instruments (the Sargan test equivalent). It is noteworthy that they confirm the validity of the six instruments described below.

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<sup>8</sup> Relaxing our main identification assumption and performing individual panel estimations (pooled OLS, fixed- and random-effects) rather than joint ones over the entire sample of countries does not alter our main conclusion: both vulnerability and trauma effects hold. These estimates are available from the authors upon request.

We start by instrumenting both endogenous variables together (column 1) and then we instrument each of them separately (columns 2 and 3). For parsimony, we remove the contemporaneous terms of each endogenous variable that are not significant (see previous subsection). We instrument NPL by the Composite Indicator of Systemic Stress (CISS), stock market volatility and the Saint Louis Fed Financial Stress Index (STLFSI) (columns 1 and 3), while we instrument credit to GDP by assets to GDP, turnover ratio and market capitalisation (columns 1 and 2).<sup>9</sup> NPL are shown to be influenced by macroeconomic and bank-specific factors like the ‘too-big-to-fail’ presumption (Louzis et al., 2012). A model of non-performing loan determination would then also include an index of systemic risk, a volatility index or an index of financial stress. Similarly, the theoretical model of the degree of credit would nest the demand side of the credit market and also draw on the supply side, hence on the liquidity and depth of the financial system. These unobservable structural characteristics are proxied by assets to GDP, turnover ratio or market capitalisation. While our instruments are not highly correlated, the consistency of the estimated results across the 3 different instruments for each instrumented variable supports the validity of the instrumental variable approach to estimate causal effects of credit or NPL one on the other.

Results of estimations with SUR-IV are reported in Table 2. They point to robust interrelationships between credit and banking fragility and to robust correlations to macro control variables, GDP growth in the equation of credit to GDP and GDP growth and inflation in the equation of NPL. In this latter equation, the correlations to the long-term interest rate and to taxes on business are also robust. There is a negative causal impact of NPL on credit to GDP and a positive causal impact of credit to GDP on NPL, suggesting that the trauma and vulnerability effects put forward in the previous section are indeed at work.

While confirming the previous estimates, both effects are of higher magnitude with 3SLS than with a SUR model only. Since our baseline results are robust to IV estimation, the rest of the analysis is performed with the SUR model so as to provide the most conservative results, i.e. with lower bound estimates rather than upper bound ones.

### *5.3. Discussion on sub-samples and different controls*

SUR estimates for subgroups of countries (Table 3) confirm the trauma effect for the Eurozone, and EU core and periphery countries; the effect is more than four times higher in core than periphery countries. Interestingly, there is a divergence for the vulnerability effect between the Eurozone and core countries on one side and periphery countries and newcomers on the other side: credit has no incidence on banking fragility in the latter. This may proceed from different stages of credit development between the core and the periphery of the EU and shed light on the threshold impact of credit to GDP ratios on banking fragility discussed in section 5.1.

The coefficients associated to the lagged values of the dependent variables are in all cases very significant and account for the persistence of these processes. We also show in Table E in the Appendix that long-term real interest rates have no impact on credit to GDP and a

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<sup>9</sup> The CISS includes 15 raw measures, mainly of market-based financial stress, which are split equally into five categories, namely the financial intermediaries sector, money markets, equity markets, bond markets and foreign exchange markets. The CISS places relatively more weight on situations in which stress prevails simultaneously in several market segments. It is unit-free and constrained to lie within the unit interval (see Hollo et al., 2012). The STLFSI is constructed on US data, but because financial markets are much integrated, at least much more than labour, goods or credit markets, we assume that this index could act as another relevant proxy for instability on financial markets in Europe.

positive impact on NPL. One possible interpretation of the coefficient associated to long-term real interest rates may be that long-term real interest rates have positive effects on the supply side of credits that offset their negative effects on the demand side. This would explain the absence of an impact on the credit to GDP ratio. High interest rates would reveal the fragility of the weakest debtors, increase the share of NPL and trigger banking fragility. The substitution effect between bank intermediation and financial markets does not appear in the data: stock market capitalisation has no significant impact on credit. In addition, stock market capitalisation has no effect on NPL. It appears that direct taxes on business are negatively correlated to banking fragility. Finally, the index of financial reform is neither correlated with credit nor with banking fragility. This is consistent with Tressel and Detragiache (2008). We find evidence that the GDP growth rate is negatively correlated to the credit to GDP ratio and to NPL. The former result might be related to different degrees of credit development in the EU and might therefore be related to the convergence effect: most developed economies in the EU share the most developed banking and financial systems; hence, these developed countries with relatively low GDP growth rates would show a more dynamic credit, whereas least-developed ones would have a less dynamic one. The negative impact of the growth rate on NPL would also match the argument of the convergence effect: the pace of growth in the least-developed-least financialised countries would not produce the same increase in risk-taking by banks and on financial markets as in the most-developed-most-financialised economies. When credit rises, the smaller economic growth rate would be synonymous of more risks, generating a rise in NPL. Evidence on the positive impact of inflation on banking fragility is strong. Finally, trade openness is not correlated to credit to GDP or banking fragility.<sup>10</sup>

#### **5.4. Introducing government debt**

We enlarge, in Table 4, the scope of common determinants of credit and banking fragility to government debt following Cooper and Nikolov (2013). First, our previous results about the vulnerability effect still hold. Second, it appears that public debt to GDP ratios have a positive effect on banking fragility in the EZ and core EU countries.<sup>11</sup> However, if we decompose this effect into normal times and crisis times, it seems that government debt impinges on banking fragility during crisis whereas the effect is null (EZ and core EU countries) or even negative (all countries or periphery EU countries) in normal times. This is consistent with the analysis of Caruana and Avdjiev (2012) and with the home bias in periphery countries that Acharya and Steffen (2015) reveal. A growing debt sustained by a home bias may reduce international financial contagion risks. Meanwhile, the trauma effect is no longer statistically significant in the Eurozone and EU core countries, and public debt to GDP ratios are negatively correlated to credit except in periphery EU countries. This supports the argument of a possible direct crowding-out effect in the core or of an indirect one in the periphery through the positive effect of higher public debt on banking fragility which may push banks to reduce their supply of credits and to deleverage.

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<sup>10</sup> This result is confirmed when replacing trade openness by an index measuring countries' degree of capital account openness, defined by Chinn and Ito (2006).

<sup>11</sup> For simplicity, we only present results for all countries, EZ, core and periphery countries. Results for core 2 and Newcomers are available upon request. Sub-sample choices do not affect our main results.

### *5.5. The stock market view of financialisation*

So far, we have focused on intermediated finance through credit. We complement the analysis by looking at direct finance through stock markets. In the EU the two types of funding are not substitutes. Because of a selection bias, households and small and mid-sized corporations do not have the same access to financial markets as large corporations. Consistent with Beck and Levine (2004), we measure financial deepening by the turnover ratio which proxies the depth and liquidity of stock markets. In parallel, financial instability is captured by stock market volatility.

Table 5 reports the estimates with this new set of variables. The opposite effects between banking fragility (now financial instability) and credit (now turnover ratio) are still captured with some subsample limitations though. On the one hand, the turnover ratio positively affects stock market volatility, except in core EU countries. This suggests that, except for the EU core, the vulnerability effect is not contingent on the definition of financialisation, whether it depends on banks or on financial markets. On the other hand, stock market volatility has a negative effect on the depth and liquidity of financial markets (the turnover ratio) in the EU core only, confirming there a trauma effect. The specificity of the EU core results may stem from its high level of financial development.

## **6. Conclusion**

We represent credit and banking fragility as a system of simultaneous joint data generating processes (estimated with Seemingly Unrelated Regressions) whose error terms are correlated and find that credit positively affects banking fragility –the vulnerability effect– and banking fragility negatively affects credit –the trauma effect–. We find evidence of some non-linearities between the two variables. NPL have a non-linear effect on credit to GDP depending on the level of credit to GDP, while the effect of credit to GDP on NPL –the vulnerability effect– depends on the level of credit to GDP and is time contingent: this effect kicks in during crisis times. In addition, we show that the existence of vulnerability and trauma effects are not exclusively related to a credit view of financialisation. Endorsing a market view of financialisation gives similar outcomes, except for the EU core: a positive effect of financial deepening –measured by the turnover ratio– on financial instability –measured by stock market volatility– and a negative effect of stock market volatility on the turnover ratio.

The existence of a vulnerability effect in the EU as a whole, in the Eurozone, in the core of the EU but not at its periphery, and of a trauma effect in all samples raises some policy recommendations. First, the existence of both effects confirms the requirement to control and supervise credit supply in the Eurozone and core countries of the EU. According to our results, monitoring credit, via policies which remain to be discussed –e.g. a change in capital adequacy ratios–, would alleviate the risks of banking fragility. Second, in the EU periphery countries, the variations in long-term interest rates and inflation play a strong role in the rise of banking fragility: hence, supervising credit dynamics in the periphery, within the Banking union, should be complemented with macroeconomic policies aimed at achieving low and stable inflation and long-term interest rates.

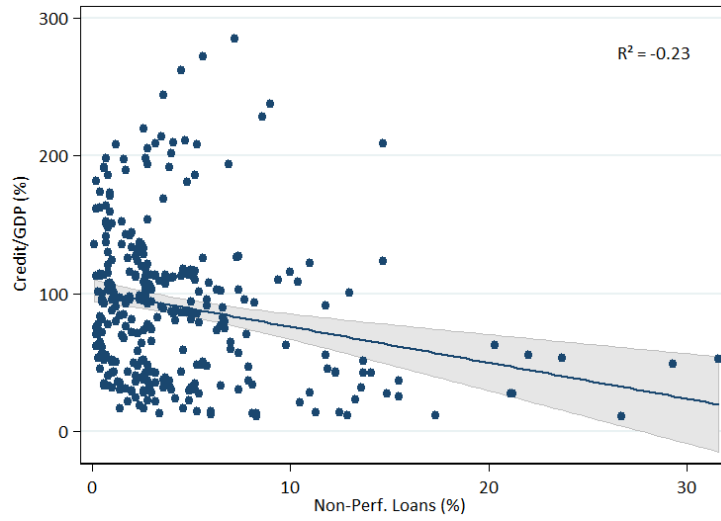
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**Figure 1 -Credit and banking fragility (Source: GFDD)**





**Table 1: Benchmark**

	(1)	(2)	(3)	(4)	(5)	(6)	(7)
	Redux	Benchmark	Square	Interaction	Crisis	3-var	Credit growth
	Non-Perf L.	Non-Perf L.	Non-Perf L.	Non-Perf L.	Non-Perf L.	Non-Perf L.	Non-Perf L.
Lag Dep. Var.	0.75*** [0.05]	0.70*** [0.05]	0.60*** [0.05]	0.70*** [0.05]	0.61*** [0.05]	0.65*** [0.05]	0.79*** [0.05]
Credit/GDP	0.38*** [0.10]	-0.02 [0.11]	-0.07 [0.10]	0.00 [0.11]	-0.06 [0.10]	0.05 [0.11]	-0.05 [0.05]
Credit/GDP <sub>t-1</sub>	-0.03 [0.11]	0.24** [0.11]	0.34*** [0.11]	0.22* [0.11]	0.12 [0.11]	0.15 [0.11]	0.20*** [0.05]
(Credit/GDP <sub>t-1</sub> ) <sup>2</sup>			0.22*** [0.05]				
Interaction				0.03 [0.04]			
Credit/GDP <sub>t-1</sub> * Crisis					0.52*** [0.11]		
Crisis					-0.01 [0.13]		
ΣCredit/GDP <sub>(t + t-1)</sub>	0.35*** [0.06]	0.22*** [0.07]	0.27*** [0.06]	0.22*** [0.07]	0.06 [0.07]	0.20*** [0.07]	0.15*** [0.05]
	Credit/GDP	Credit/GDP	Credit/GDP	Credit/GDP	Credit/GDP	Credit/GDP	Credit/GDP
Lag Dep. Var.	0.86*** [0.03]	0.86*** [0.05]	0.86*** [0.05]	0.86*** [0.05]	0.85*** [0.05]	0.86*** [0.05]	0.37*** [0.06]
Non-Perf L.	0.12*** [0.03]	-0.01 [0.05]	0.00 [0.05]	0.00 [0.05]	0.00 [0.05]	0.02 [0.05]	-0.10 [0.09]
Non-Perf L <sub>t-1</sub>	-0.25*** [0.03]	-0.14*** [0.05]	-0.16*** [0.05]	-0.14*** [0.05]	-0.11** [0.05]	-0.13*** [0.05]	-0.44*** [0.09]
(Non-Perf L <sub>t-1</sub> ) <sup>2</sup>			0.03 [0.03]				
Interaction				-0.07** [0.03]			
Non-Perf L <sub>t-1</sub> * Crisis					-0.06 [0.07]		
Crisis					0.11 [0.11]		
ΣNon-Perf L <sub>(t + t-1)</sub>	-0.12*** [0.03]	-0.15*** [0.03]	-0.16*** [0.04]	-0.14*** [0.03]	-0.12** [0.05]	-0.11*** [0.04]	-0.54*** [0.06]
Controls X <sub>it</sub>	No	Yes	Yes	Yes	Yes	Yes	No
Controls Z <sub>it</sub>	Yes	Yes	Yes	Yes	Yes	Yes	Yes
3-equation model	No	No	No	No	No	Yes	No
N	275	182	182	182	182	179	253
R <sup>2</sup> <sub>1</sub>	0.61	0.75	0.78	0.75	0.78	0.74	0.60
R <sup>2</sup> <sub>2</sub>	0.89	0.89	0.89	0.89	0.89	0.89	0.39

Standard errors in brackets. \* p < 0.1 , \*\* p < 0.05, \*\*\* p < 0.01. Estimated from equation (1). All variables are standardized to a normal distribution by country. The interaction term is between the lag of the dependent variable and credit/GDP in the upper panel, and non-performing loans in the lower panel. In column (6), the SUR model is estimated with 3 dependent variables: non-performing loans, credit/GDP, and long-term interest rates, and the overall model is augmented with short-term interest rates. For sake of simplicity, the 3rd equation for long-term interest rates and the parameters for short-term interest rate are not shown here. They are available from the authors upon request. In column (7), the credit variable, in level, is replaced by the credit growth.

**Table 2: SUR-IV 3SLS estimation**

	(1)	(2)	(3)
	All	All	All
Instrumented	Credit/GDP Non-Perf L.	Credit/GDP	Non-Perf L.
Instruments	CISS / Asset/GDP Volat / Turnover STLFSI / Market Cap.	Asset/GDP Turnover Market Cap.	CISS Volat STLFSI
	Non-Perf L.	Non-Perf L.	Non-Perf L.
Lag Dep. Var.	0.70*** [0.04]	0.70*** [0.04]	0.70*** [0.04]
Credit/GDP <sub>t-1</sub>	0.22*** [0.06]	0.22*** [0.06]	0.22*** [0.06]
Regression of 3SLS residuals on instruments			
R <sup>2</sup>	0.12	0.03	0.06
	Credit/GDP	Credit/GDP	Credit/GDP
Lag Dep. Var.	0.86*** [0.04]	0.86*** [0.04]	0.86*** [0.04]
Non-Perf L <sub>t-1</sub>	-0.15*** [0.03]	-0.15*** [0.03]	-0.15*** [0.03]
Regression of 3SLS residuals on instruments			
R <sup>2</sup>	0.08	0.03	0.06
Controls X <sub>it</sub>	Yes	Yes	Yes
Controls Z <sub>it</sub>	Yes	Yes	Yes
N	182	182	182

Standard errors in brackets. \* p < 0.1, \*\* p < 0.05, \*\*\* p < 0.01. Estimated from equation (1). All variables are standardized to a normal distribution by country.

**Table 3: Geographical zones**

	(1) All	(2) EZ	(3) Core	(4) Core 2	(5) Periphery	(6) Newcomers
	Non-Perf L.	Non-Perf L.	Non-Perf L.	Non-Perf L.	Non-Perf L.	Non-Perf L.
Lag Dep. Var.	0.70*** [0.05]	0.69*** [0.06]	0.58*** [0.07]	0.66*** [0.05]	0.75*** [0.07]	0.50*** [0.15]
Credit/GDP	-0.02 [0.11]	-0.10 [0.15]	-0.44*** [0.13]	-0.24** [0.12]	0.43** [0.20]	1.03*** [0.40]
Credit/GDP <sub>t-1</sub>	0.24** [0.11]	0.40** [0.16]	0.66*** [0.13]	0.48*** [0.12]	-0.31 [0.22]	-0.97** [0.41]
$\Sigma$ Credit/GDP <sub>(t+t-1)</sub>	0.22*** [0.07]	0.30*** [0.08]	0.22*** [0.08]	0.24*** [0.07]	0.13 [0.12]	0.07 [0.11]
	Credit/GDP	Credit/GDP	Credit/GDP	Credit/GDP	Credit/GDP	Credit/GDP
Lag Dep. Var.	0.86*** [0.05]	0.94*** [0.05]	0.94*** [0.06]	0.88*** [0.06]	0.89*** [0.06]	0.98*** [0.05]
Non-Perf L.	-0.01 [0.05]	-0.03 [0.05]	-0.28*** [0.08]	-0.14** [0.07]	0.11** [0.05]	0.21*** [0.08]
Non-Perf L <sub>t-1</sub>	-0.14*** [0.05]	-0.10** [0.05]	0.02 [0.07]	-0.03 [0.06]	-0.18*** [0.05]	-0.32*** [0.05]
$\Sigma$ Non-Perf L <sub>(t+t-1)</sub>	-0.15*** [0.03]	-0.13*** [0.03]	-0.26*** [0.06]	-0.17*** [0.04]	-0.06* [0.04]	-0.11 [0.08]
Controls X <sub>it</sub>	Yes	Yes	Yes	Yes	Yes	Yes
Controls Z <sub>it</sub>	Yes	Yes	Yes	Yes	Yes	Yes
N	182	126	92	118	90	27
R <sup>2</sup> <sub>1</sub>	0.75	0.78	0.74	0.76	0.82	0.90
R <sup>2</sup> <sub>2</sub>	0.89	0.92	0.86	0.88	0.95	0.98

Standard errors in brackets. \* p < 0.1, \*\* p < 0.05, \*\*\* p < 0.01. Estimated from equation (1). All variables are standardized to a normal distribution by country. The composition of country groups is presented in Table C in the Appendix.

**Table 4: Introducing government debt**

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	All	EZ	Core	Periphery	All	EZ	Core	Periphery
	Non-Perf L.	Non-Perf L.	Non-Perf L.	Non-Perf L.	Non-Perf L.	Non-Perf L.	Non-Perf L.	Non-Perf L.
Lag Dep. Var.	0.03 [0.11]	0.03 [0.15]	-0.34** [0.13]	0.46** [0.20]	-0.02 [0.10]	-0.09 [0.15]	-0.33*** [0.12]	0.14 [0.21]
Credit/GDP	0.20* [0.11]	0.30* [0.16]	0.57*** [0.14]	-0.37 [0.23]	0.21** [0.10]	0.37** [0.16]	0.54*** [0.13]	-0.17 [0.21]
Credit/GDP <sub>t-1</sub>	0.68*** [0.05]	0.66*** [0.06]	0.57*** [0.07]	0.77*** [0.07]	0.65*** [0.05]	0.67*** [0.06]	0.55*** [0.07]	0.71*** [0.07]
Gov. Debt	0.08 [0.06]	0.13* [0.07]	0.07 [0.07]	-0.09 [0.10]	-0.22*** [0.08]	-0.17 [0.14]	-0.20* [0.10]	-0.48*** [0.13]
Gov. Debt * Crisis					0.47*** [0.10]	0.38** [0.15]	0.44*** [0.13]	0.63*** [0.15]
Crisis					0.36*** [0.14]	0.33 [0.20]	0.37* [0.21]	0.43** [0.19]
$\Sigma$ Credit/GDP <sub>(t + t-1)</sub>	0.24*** [0.07]	0.33*** [0.08]	0.23*** [0.08]	0.09 [0.13]	0.19*** [0.06]	0.28*** [0.08]	0.21*** [0.07]	-0.03 [0.12]
	Credit/GDP	Credit/GDP	Credit/GDP	Credit/GDP	Credit/GDP	Credit/GDP	Credit/GDP	Credit/GDP
Lag Dep. Var.	0.84*** [0.05]	0.91*** [0.05]	0.92*** [0.06]	0.92*** [0.06]	0.83*** [0.05]	0.90*** [0.05]	0.92*** [0.06]	0.87*** [0.06]
Non-Perf L.	0.02 [0.05]	0.01 [0.05]	-0.20** [0.08]	0.12** [0.05]	-0.01 [0.05]	-0.03 [0.05]	-0.23*** [0.08]	0.04 [0.05]
Non-Perf L <sub>t-1</sub>	-0.13*** [0.05]	-0.09** [0.05]	0.02 [0.07]	-0.19*** [0.05]	-0.10** [0.05]	-0.06 [0.05]	0.03 [0.07]	-0.12** [0.05]
Gov. Debt	-0.09** [0.04]	-0.10** [0.04]	-0.13** [0.05]	0.06 [0.05]	-0.12** [0.06]	-0.22*** [0.08]	-0.19** [0.08]	0 [0.07]
Gov. Debt * Crisis					0.03 [0.07]	0.15* [0.09]	0.08 [0.11]	0.06 [0.08]
Crisis					0.19* [0.10]	0.22* [0.12]	0.18 [0.18]	0.36*** [0.09]
$\Sigma$ Non-Perf L <sub>(t + t-1)</sub>	-0.11*** [0.04]	-0.09** [0.04]	-0.19*** [0.06]	-0.07* [0.04]	-0.11*** [0.04]	-0.09** [0.04]	-0.19*** [0.06]	-0.09** [0.04]
Controls X <sub>i,t</sub>	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Controls Z <sub>i,t</sub>	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
N	182	126	92	90	182	126	92	90
R <sup>2</sup> <sub>1</sub>	0.75	0.78	0.75	0.82	0.78	0.79	0.78	0.85
R <sup>2</sup> <sub>2</sub>	0.89	0.93	0.87	0.95	0.89	0.93	0.87	0.96

Standard errors in brackets. \* p < 0.1, \*\* p < 0.05, \*\*\* p < 0.01. Estimated from equation (1). All variables are standardized to a normal distribution by country.

**Table 5: Stock market view of financialisation**

	(1) All	(2) EZ	(3) Core	(4) Periphery
	Volat	Volat	Volat	Volat
Lag Dep. Var.	0.53*** [0.04]	0.55*** [0.05]	0.50*** [0.07]	0.51*** [0.06]
Turnover	0.05 [0.05]	0.07 [0.06]	-0.12 [0.09]	0.16*** [0.06]
Turnover <sub>t-1</sub>	0.23*** [0.05]	0.21*** [0.06]	0.15 [0.10]	0.32*** [0.06]
$\Sigma$ Turnover <sub>(t + t-1)</sub>	0.28*** [0.05]	0.27*** [0.06]	0.03 [0.10]	0.48*** [0.07]
	Turnover	Turnover	Turnover	Turnover
Lag Dep. Var.	0.48*** [0.07]	0.45*** [0.08]	0.47*** [0.10]	0.22* [0.12]
Volat	0.08 [0.09]	0.12 [0.12]	-0.15 [0.11]	0.44*** [0.17]
Volat <sub>t-1</sub>	-0.17** [0.08]	-0.21** [0.09]	-0.1 [0.09]	-0.24* [0.13]
$\Sigma$ Volat <sub>(t + t-1)</sub>	-0.08 [0.07]	-0.09 [0.09]	-0.25*** [0.09]	0.20 [0.13]
Controls $X_{i,t}$	Yes	Yes	Yes	Yes
Controls $Z_{i,t}$	Yes	Yes	Yes	Yes
N	200	138	107	93
R <sup>2</sup> <sub>1</sub>	0.67	0.71	0.68	0.75
R <sup>2</sup> <sub>2</sub>	0.42	0.39	0.58	0.34

Standard errors in brackets. \* p < 0.1, \*\* p < 0.05, \*\*\* p < 0.01. Estimated from equation (1). All variables are standardized to a normal distribution by country.

## APPENDIX

**Table A: Data Description and Sources**

Abbreviation	Description	Source	Frequency
Credit/GDP	Private credit by deposit money banks and other financial institutions to GDP (%)	GFDD	annual
Non-Perf L.	Bank non-performing loans to gross loans (%)	GFDD	annual
Asset/GDP	Deposit money banks' assets to GDP (%)	GFDD	annual
Turnover	Stock market turnover ratio (%)	GFDD	annual
CISS (composite indicator of systemic stress)	Index comprising the five most important segments of a financial system: bank and non-bank financial intermediaries sector, money markets, securities markets and foreign exchange markets.	ECB	weekly aggregated to annual
STLFSI	St. Louis Fed Financial Stress Index	FRED	annual
Volat	Stock price volatility (%)	GFDD	annual
LT Real IR	Real long term interest rates (difference between long term interest rates and inflation)	Authors calculation using OECD & WDI	annual
Market Cap.	Market capitalisation of listed companies (% of GDP)	WDI	annual
Tax. Business	Cyclically adjusted direct taxes on business (% of GDP)	OECD	annual
Gov. Debt	Gross public debt, Maastricht criterion, as % of GDP	OECD	annual
Fin. Reform	Index of financial reform	IMF	annual
Inflation	Inflation, consumer prices (annual %)	WDI	annual
GDP growth	GDP growth (annual %)	WDI	annual
Trade Open.	Trade (% of GDP)	WDI	annual

**Table B: Descriptive statistics**

Variable	Obs	Mean	Std. Dev.	Min	Max
Main variables					
Credit/GDP	344	93.12	57.61	6.38	284.62
Non-Perf L.	343	4.75	5.01	0.10	31.60
Financial controls					
LT Real IR	277	2.30	2.03	-1.72	21.00
Market Cap.	405	53.80	47.05	2.41	323.66
Tax. Business	278	0.21	0.55	0.01	3.44
Fin. Reform	330	0.92	0.08	0.49	1.00
Macro controls					
Inflation	405	3.68	5.16	-4.48	59.10
GDP growth	405	2.55	3.68	-17.95	12.23
Trade Open.	397	110.09	52.52	46.64	333.53

**Table C: Subsamples composition**

Eurozone (EZ)	Core	Core 2	Newcomers	Periphery
Austria	Austria	Austria	Bulgaria	Bulgaria
Belgium	Belgium	Belgium	Cyprus	Cyprus
Germany	Germany	Germany	Czech Republic	Czech Republic
Spain	Denmark	Denmark	Estonia	Estonia
Finland	Finland	Finland	Hungary	Spain
France	France	France	Lithuania	Greece
Greece	Luxembourg	Luxembourg	Latvia	Hungary
Ireland	Netherlands	Netherlands	Malta	Ireland
Italy	Sweden	Sweden	Poland	Italy
Netherlands	United Kingdom	United Kingdom	Romania	Lithuania
Portugal		Italy	Slovenia	Latvia
		Spain	Slovakia	Malta
				Poland
				Portugal
				Romania
				Slovenia
				Slovakia

**Table D: Mean of the main variables for the different subsamples**

	All		Core		Core 2		Periphery		Newcomers	
	Mean	SD	Mean	SD	Mean	SD	Mean	SD	Mean	SD
Non-Perf L. (%)	4.78	3.01	2.19	1.22	2.68	2.22	6.30	2.69	6.95	6.32
Credit/GDP (% of GDP)	91.35	50.87	116.01	30.12	115.34	41.34	76.85	54.90	62.93	59.06
LT Real IR	2.28	0.58	2.16	0.30	2.15	1.31	2.42	0.76	2.09	1.67
Market Cap. (% of GDP)	53.80	40.65	91.05	39.90	86.05	49.74	31.89	19.83	22.32	17.91
Tax. Business (% of GDP)	0.20	0.53	0.08	0.11	0.07	0.10	0.31	0.70	0.65	0.98
Fin. Reform (index)	0.92	0.07	0.95	0.06	0.95	0.06	0.90	0.08	0.89	0.10
Inflation (annual %)	3.68	3.32	1.90	0.30	1.86	2.54	4.72	3.81	3.37	4.33
GDP growth (annual %)	2.55	1.13	1.96	0.65	2.00	0.96	2.91	1.20	5.57	7.21
Trade Open. (% of GDP)	110.40	50.76	112.37	66.18	102.85	65.17	109.24	39.03	120.98	32.78

**Table E: Coefficients for controls in Table 1**

	(2)	(3)	(4)	(5)	(6)
	All	All	All	All	All
	Non-Perf L.	Non-Perf L.	Non-Perf L.	Non-Perf L.	Non-Perf L.
LT Real IR	0.19*** [0.06]	0.16*** [0.06]	0.18*** [0.07]	0.21*** [0.06]	0.41*** [0.07]
Market Cap.	0.03 [0.05]	0.03 [0.04]	0.03 [0.05]	0.04 [0.04]	0.05 [0.05]
Tax. Business	-0.10** [0.04]	-0.05 [0.04]	-0.09** [0.04]	-0.08** [0.04]	-0.09** [0.04]
Fin. Reform	-0.09 [0.50]	-0.21 [0.48]	-0.12 [0.50]	-0.38 [0.48]	0.05 [0.50]
GDP growth	-0.29*** [0.05]	-0.28*** [0.05]	-0.30*** [0.06]	-0.28*** [0.05]	-0.26*** [0.06]
Inflation	0.15** [0.06]	0.10* [0.06]	0.14** [0.06]	0.14*** [0.06]	0.32*** [0.07]
Trade Open.	-0.04 [0.05]	-0.06 [0.05]	-0.03 [0.05]	-0.04 [0.05]	0.00 [0.06]
	Credit/GDP	Credit/GDP	Credit/GDP	Credit/GDP	Credit/GDP
LT Real IR	0.01 [0.05]	0.01 [0.05]	0.02 [0.05]	0.00 [0.05]	-0.06 [0.05]
Market Cap.	0.03 [0.03]	0.03 [0.03]	0.03 [0.03]	0.03 [0.03]	0.03 [0.03]
Tax. Business	0.02 [0.03]	0.01 [0.03]	0.01 [0.03]	0.01 [0.03]	0.01 [0.03]
Fin. Reform	-0.46 [0.35]	-0.43 [0.35]	-0.36 [0.35]	-0.37 [0.35]	-0.41 [0.35]
GDP growth	-0.14*** [0.04]	-0.13*** [0.04]	-0.12*** [0.04]	-0.12*** [0.04]	-0.12*** [0.04]
Inflation	0.00 [0.04]	-0.01 [0.04]	0.01 [0.04]	-0.01 [0.04]	-0.08* [0.05]
Trade Open.	0.03 [0.04]	0.02 [0.04]	0.02 [0.04]	0.01 [0.04]	0 [0.04]
Controls $X_{i,t}$	Yes	Yes	Yes	Yes	Yes
Controls $Z_{i,t}$	Yes	Yes	Yes	Yes	Yes
3-equation model	No	No	No	No	Yes
N	182	182	182	182	179
R <sup>2</sup> _1	0.75	0.78	0.75	0.78	0.74
R <sup>2</sup> _2	0.89	0.89	0.89	0.89	0.89

Standard errors in brackets. \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ . Estimated from equation (1). All variables are standardized to a normal distribution by country. For sake of simplicity, the 3rd equation for long-term interest rates and the parameters for short-term interest rate are not shown here. They are available from the authors upon request.